# Construction of Risk-Averse Enhanced Index Funds 

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#### Abstract

We propose a partial replication strategy to construct risk-averse enhanced index funds. Our model takes into account the parameter estimation risk by defining the asset returns and the return covariance terms as random variables. The variance of the index fund return is forced to be below a low-risk threshold with a large probability, thereby limiting the market risk exposure of the investors and the moral hazard associated with the wage structure of fund managers. The resulting stochastic integer problem is reformulated through the derivation of a deterministic equivalent for the risk constraint and the use of a block decomposition technique. We develop an exact outer approximation method based on the relaxation of some binary restrictions and the reformulation of the cardinality constraint. The method provides a hierarchical organization of the computations with expanding sets of integer-restricted variables and outperforms the Bonmin and the Cplex 12.1 solvers. The method can solve very large (up to 1000 securities) instances, converges fast, scales well, and is general enough to be applicable to problems with buy-in threshold constraints. Cross-validation tests show that the constructed funds track closely and are consistently less risky than the benchmark on the out-of-sample period.


## 1 Introduction

The objective of an index tracking strategy is to create an index fund whose performance replicates, as closely as possible, that of a financial market benchmark (such as the Standard and Poor's 500 Index or the Goldman Sachs Commodity Index [10]). Index tracking is a passive, also called buy-and-hold, investment strategy. Index funds undergo very limited rebalancing operations, resulting in minimal transaction costs, trading commissions and low management fees. This contrasts with active investment strategies in which the fund managers constantly rebalance their portfolio assets in an attempt to beat the market In this paper, we develop an approach to construct enhanced index funds (EIF), also called "index-plus-alpha" [53] or "alpha tilt" [18] index funds. Enhanced indexation is a structured investment approach that builds on the strengths of traditional index investing and is aimed at outperforming it [18]. It combines passive and active management techniques [3, 81] and can be viewed as a way to eschew the passive index tracking approach in favor of a semi-active approach closer to portfolio management. It resembles passive management, since it essentially uses index-oriented investment strategies which do not typically allow managers to construct funds that significantly deviate from the benchmark [62]. By contrast to traditional indexation, EIF managers are allowed to engage in limited (risk-controlled) active strategies which offer (adjusted) return enhancements relative to the benchmark return [48]. Loftus differentiates traditional from enhanced index funds in terms of the tracking error, expected alpha and information ratio [60].

Index funds are very attractive for institutional and corporate investors as well as for individuals [38]. For example, corporate pension funds are reportedly investing more than $25 \%$ of their equity holdings in index funds and about $30 \%$ of the US households having mutual funds own an index mutual fund [44]. There has been a steady increase in the proportion of capital invested in index funds since 1995 [63]. The Financial Research Corporation notes that, while $19 \%$ of the money invested in mutual funds in 1998 were directed to index funds, this proportion increased to $45 \%$ in 1999. The volume of enhanced index assets grew ten-fold between 1994 and 2000 [48] and represented more than $20 \%$ of the total indexed assets in 2000-2003 [82, 90]. In July 2009, forty-eight of the largest US financial firms reported to have $\$ 217.3$ billion in US international tax-exempt assets

[^0]under internal enhanced index management [54]. Several exchange-traded funds (ETF), such as the PowerShares FTSE RAFI U.S. Portfolio fund, are based on enhanced indexation strategies.

The growing popularity of indexation strategies lies in their ability to generate an attractive return level, outperforming active investment strategies, while exposing the investor to limited risks and low operating and management expenses [22, 64]. Considering a sample of 355 equity mutual funds active in 1970, Malkiel shows that only 158 of them have survived until 2001, and that only five of these have generated returns that were two basis points or more higher than the returns provided by index funds [64]. In the same vein, Siegel reports that the average (over the period 1974-2004) annual return of all actively-managed mutual funds trailed the S\&P 500 Index and the Wilshire 5000 Index by, respectively, 87 and 105 basis points [86]. Siegel advocates that an index fund whose average performance is identical to that of the market index outperforms most actively-managed portfolios. The best EIFs have reportedly outperformed their traditional index counterparts by $1 \%$ to $3 \%$ per year [88]. Since traditional index funds outperform some of the active funds by a similar margin over the long term, enhanced indexation versus active investing can lead to a difference of $2 \%$ to $6 \%$ per year relative to the majority of active investing approaches. The "superior" total return performance is usually attributed to the semi-active management of EIFs which allows taking advantage from rapidly changing market conditions [82]. Although lower than those of most mutual funds, EIFs' expense ratios and turnover rates are however higher than those of traditional index funds. As a basis for comparison, the expense ratios of the index funds tracking the S\&P 500 and the Vanguard 500 indices amount to $0.2 \%$ and $0.18 \%$ respectively. On the other hand, EIFs' expense ratios are typically between $0.3 \%$ and $0.7 \%$ (in 2002, Jorion reported them to average $0.32 \%$ [48]), while they range from $1 \%$ to $1.5 \%$ for actively-managed funds. Enhanced indexation also exposes investors to a higher risk than traditional indexation methods. Traditional index funds are only exposed to the market risk stemming from the volatility of the market. On the other hand, EIFs are semi-actively managed and thus expose investors to management risk, i.e., the risk associated with ineffective active fund management.

Index funds can be built with a full or partial replication approach. The full replication approach consists in buying all assets included in the financial index in the same proportions as in the index [22]. The full replication approach leads to frequent rebalancing of the portfolio, high transaction costs, and forces managers to hold nonliquid positions. Moreover, managers cannot trade the stocks at their fair prices. Indeed, arbitrageurs use the lag between the index announcement and the fund rebalancing to take positions on stocks entering and leaving the index, which causes an artificial inflation or deflation of stock prices. For instance, the price to pay for the full replication of the Russell 2000 index is estimated at $1.3 \%$ and $1.84 \%$ annually. These issues, well-documented in $[5,12,18,28,29,31]$, hamper the use of a full replication approach and explain the success of partial replication approaches. Partial replication means that the fund manager is allowed to invest in a limited number of securities to track the benchmark $[5,8,28,38,66,69,80]$. The requirement is enforced through the use of binary decision variables and the introduction of a cardinality constraint. The transaction and administration costs are typically defined as an increasing function of the number of assets in the portfolio [25]. By limiting the number of assets that can be included in the index fund, the cardinality constraint actually enforces an upper bound on the above costs. The limitation of these costs is especially important when the portfolio has a small net asset value [46].

Index models based on partial replication are NP-hard and pose severe computational challenges [25]. Multiple models and algorithmic techniques have been proposed for their construction. A comprehensive review of the literature (until 2003) can be found in [5]. Below, we review a number of more recent index tracking studies, starting with traditional index funds before moving to the enhanced indexation literature. We distinguish the studies in terms of the objective function, the dimension of the asset universe, and the type of solution method.

An evolutionary heuristic [5], a clustering approach [37], and a Lagrangian relaxation method used within a branch-and-bound algorithm [85] are employed to construct an index fund from an asset universe containing 225 (in [5]) and 500 (in [85]) securities. The mean squared tracking error is minimized and asset universes comprising respectively $225,65,30$ and 225 assets are analyzed in [55, 66, 76, 80]. The index fund is built with a differential evolution search heuristic in [55, 66]. A genetic algorithm determines the amount to be invested in the assets included in the index fund in [76], while Ruiz-Torrubiano and Suarez develop a genetic algorithm to select the assets and solve a quadratic programming problem to define the size of the positions [80]. Two weighted components representing the tracking error variance and the number of assets in the index are included in the objective function in [25, 46]. The solution method is based on a continuous approximation of the discrete function. The tracking error variance is minimized with a set of local heuristics integrated in a decision support system in [32]. The absolute value of the return tracking error is minimized in [8, 87]. In [87], a decomposition method is proposed to solve the two-stage programming problem. Four tracking error functions are analyzed in [38]. The associated quadratic integer problems are solved for an asset universe comprising 65 assets using a tailored heuristic approach. Yao et al. [93] use a control theory approach to formulate the index problem. A semi-definite programming approach is applied to asset universes containing four and five stocks.

As underlined by Canakgoz and Beasley [18], the first studies devoted to the construction of EIFs go back to 2005. In [33], the objective function is a convex combination of tracking error and excess return. A clustering method determines which of the 487 considered assets are included in the index fund. The mean absolute deviation of the difference between the benchmark return and that of the constructed index fund increased by a positive constant is minimized in [53]. The construction of the EIF accounts for transaction costs and involves the minimization of a separable concave function under linear constraints [92]. The method is applied to a problem considering the 225 assets of the Nikkei index. In [2], two return time-series are generated by respectively adding and withdrawing a fixed positive return (alpha) from the return of the tracked index. A cointegration approach is then used to build two portfolios tracking the "alpha-plus" and the "alpha-minus" time-series. The index fund is constructed by going long on the alpha-plus portfolio and shorting on the alpha-minus one. The method requires the inclusion of a relatively large number of stocks in order to consistently reproduce the return of the benchmark. Canakgoz and Beasley formulate the construction of an EIF as a mixed-integer problem [18]. A three-step solution method involving the computation of a regression intercept and slope is proposed to build index funds that comprise between $3 \%$ and $32 \%$ of the considered assets (up to 2151). In [92], a goal programming formulation is derived to minimize the sum of the deviations from the desired levels of return and tracking error. The method is applied to a sample ( 426 stocks) of the Taiwanese stock market. ChávezBedoya and Birge develop a parametric approach for constructing traditional and enhanced index funds [22]. The proposed nonlinear formulation has a weighted multi-objective function that represents the correlation between the index and the benchmark returns, the ratio of their return standard deviations, and the average excess return of the fund over the benchmark. Setting the weight of the excess return to zero results in the construction of a standard index fund. The method is used to construct indices containing 25 to 75 assets selected from a 475 -asset universe. The out-of-sample tests show that the in-sample and out-of-sample correlation and ratio of standard deviations are close, while, on the other hand, the out-of-sample excess return of the index deviates more from its in-sample value. Jorion [48] observes that the EIFs closely tracking a benchmark exhibit a variance level that is most often higher than the one of the benchmark. He empirically shows that derivatives-based EIFs are those that have the risk profile most closely resembling that of the benchmark. Note that many of the index fund studies (see, e.g., [ $5,33,18,46,55,66,76,80,92]$ ) use heuristic methods and model the future asset returns as
deterministic parameters. The presence of a cardinality constraint and integer variables, a key feature and source of complexity for the partial replication index problem, arise in other types of financial optimization problems (see, e.g., [9, 15, 21, 47]).

The contribution of this study is the introduction of a new model and exact solution method for the construction of risk-averse enhanced index funds. From a modeling perspective, the proposed approach has four key contributions. First, the model incorporates a constraint that limits the global variance of the constructed fund and enables the pursuit of a risk-averse enhanced indexation approach. The risk-averse feature stems from the fact that the constructed EIF does not only closely track the benchmark, but does so while controlling the variance of the EIF's return, thus limiting the risk exposure of the investor. It is a critical property, since previous studies [7, 29, 48, 49, 77] have shown that it is possible to use (semi-)active allocation strategies to construct portfolios that track very closely a benchmark but that have a much larger variance than that of the benchmark [30, 68]. In a study devoted to EIFs, Jorion suggests that the minimization of the tracking error should be subjected to the satisfaction of a constraint limiting the global variance of the fund [48]. Based on the same considerations, Chow proposes to minimize a multi-objective function that includes a tracking error variance term and another one for the variance of the portfolio's return [24]. The above discussion raises the question whether the similarity between the returns of the EIF and of the market index is in line with the risk undertaken by holding the index fund. Second, the risk constraint imposes an upper bound on the EIF's return variance and thereby limits an agency problem associated with the performance-fee wage structure of fund managers [49]. Performance fees can be assimilated to an option-like pattern in the manager's salary [49]. In some instances, they provide an incentive to take on more risk to increase the value of the option [41, 49]. If the same level of tracking error can be obtained by several funds, the manager could be better off by opting for the one with the highest variance. For similar reasons, Grinblatt and Titman recommend the inclusion of covenants defining allowable portfolio strategies in performance-based contracts [40]. Third, the proposed model explicitly accounts for the volatile character of the returns of securities. Most previous index fund studies assess the return of an asset by a point-estimate. This opens the door to the estimation risk [4] and its drawbacks (see, e.g., [17, 23, 70]). In contrast to this, and to account for our incomplete knowledge of the return behavior, we model asset returns as random variables and assume that they are driven by a factor model. We propose a new stochastic risk factor model in which factor returns are also random variables, and we derive a new second-order cone formulation equivalent to the probabilistic risk constraint. Fourth, we use a decomposition method to obtain a much sparser representation of the variance-covariance matrix. The control of the market risk exposure and the taking into account of the estimation risk leads to the formulation of a stochastic integer problem which imposes that the variance of the EIF does not exceed, with a high probability, a prescribed level of return variability. From the optimization angle, we propose a new outer approximation solution approach that is highly efficient to solve the nonlinear, NP-hard formulations used for the construction of EIFs. The models include a quadratic objective function, a probabilistic constraint, a cardinality constraint and its associated integrality restrictions. Our solution method is robust, very fast in finding the optimal solution for problems including up to 1000 assets, and scales very well. These results must be put in parallel with recent studies $[55,66]$ stating that exact solution methods cannot handle the cardinality constraints present in index tracking problems. Cross-validation tests show that the EIFs track closely and are less risky than the benchmark (the S\&P 500 index fund) on the out-of-sample period. The Sharpe ratio of the vast majority of the EIFs is also higher than that of the benchmark.

The rest of the paper is organized as follows. Section 2 describes the problem formulations. In Section 3, we propose two variants of the outer approximation solution method. Section 4 presents the results of a
computational study that evaluates the performance of our approach. Concluding remarks are given in Section 5 .

## 2 Problem Formulation

In this section, we first present the stochastic integer model proposed for the construction of an EIF. Second, we describe the stochastic factor model for the asset returns. Third, we derive a new deterministic reformulation of the stochastic risk constraint and analyze the properties of the resulting deterministic problem. Fourth, we detail the block-decomposition method that provides a sparser representation of the variance-covariance matrix. Finally, we present another risk-averse EIF model that prevents the investor from holding very small positions.

### 2.1 Stochastic Index Tracking Model

Consider an asset universe composed of one riskless and $n$ risky assets which an investor can include in a fund that replicates a benchmark market index $M$. The returns of the risky assets and the return of the market index have been observed over $l$ consecutive periods. We denote by $r_{M}^{t}$ the observed return of the market index at period $t$ and by $r_{i}^{t}$ the observed return of asset $i$ at $t$. The position (i.e., proportion of capital invested) in each security is represented by the vector $\tilde{w}$ : $\tilde{w}_{0}$ is the position in the riskless asset, while $\tilde{w}_{i}, i=1, \ldots, n$ refers to the position in the risky asset $i$. The return on the risky asset $i$ at future periods is a stochastic variable denoted by $\xi_{i}$. The only probabilistic information assumed about the random asset return vector $\xi$ is that it has finite first and second moments; it has an $n$-variate distribution with mean vector $\mu$ and variance-covariance (VC) matrix $\Sigma=E\left[(\xi-\mu)(\xi-\mu)^{\prime}\right]$. Let $\tilde{\mu}=\left[\mu_{0} \mu\right]$ denote the $(n+1)$-dimensional mean return ( $\mu_{0}$ is the return of the riskless asset) and let $w$ be the $n$-dimensional vector of risky asset positions. The expected value of the return of the index fund is then $\tilde{w}^{\prime} \tilde{\mu}$ and its variance is $w^{\prime} \Sigma w$. The dimensions of the vectors and matrices are: $r_{M} \in \mathcal{R}^{l}, \xi \in \mathcal{R}^{n}, \tilde{w} \in \mathcal{R}^{n+1}, w \in \mathcal{R}^{n}$ and $r \in \mathcal{R}^{(n+1) \times l}$.

The proposed model for the construction of an index fund tracking the market index $M$ is an integer probabilistically constrained mathematical programming problem (SIP) with random technology matrix [50]:

$$
\begin{array}{cc}
(\mathbf{S I P}): & \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \\
\text { subject to } & \tilde{w}^{\prime} e=1 \\
\tilde{w} \leq \gamma \\
& \gamma^{\prime} e \leq K \\
\mathbb{P}\left(w^{\prime} \Sigma w \leq v\right) \geq p \\
& \tilde{w} \geq 0 \\
& \gamma \in\{0,1\} . \tag{7}
\end{array}
$$

The notation $e$ denotes an all-one vector. The symbol $\mathbb{P}$ represents a probability measure and $p$ and $v$ are parameters: $p$ is a probability level typically defined on $[0.7,1)$, and $v$ is the upper bound on the variance of the constructed EIF. The objective function (1) is quadratic. In order to track the performance of the market index as closely as possible, (1) minimizes the total squared deviation between the past returns of the EIF and those of the market index. The idea is that a portfolio that yielded in the past performance levels close to those of the benchmark will continue to do so in the future. The (total or average) squared deviation is one of the most popular tracking measures $[5,28,38,76]$. Other tracking metrics minimize the tracking error variance, the mean absolute deviation, the root squared mean error, and a power function of deviation (see [38] for a detailed discussion of the most popular tracking measures). The budget constraint (2) ensures that the whole available
capital is invested. The non-negativity constraint (6) precludes short-selling. Each component $\gamma_{i}$ of the binary decision vector $\gamma$ (7) indicates whether the investor holds a position in security $i$. Constraint (3) forces $\gamma_{i}$ to be equal to one if $\tilde{w}_{i}$ is strictly positive. The cardinality constraint (4) permits a partial replication strategy. It bounds from above the number $K$ of securities in which the investor can hold positions. The value of $K$ is defined to limit the administration and transaction costs which increase with the number of assets included in the index [25, 46]. In view of the difficulty to derive a precise point-estimate of the asset returns and their covariance terms, we model them as random variables. Thus, the variance $w^{\prime} \Sigma w$ of the return of the index fund is also a stochastic variable. The objective of limiting the investor's exposure to the risk entailed by the EIF is modeled using the probabilistic constraint (5). Constraint (5) ensures that the variance $w^{\prime} \Sigma w$ of the index fund return is below a low variability threshold $v$ with large (close to 1 ) probability. Hence, our model permits the construction of a risk-averse EIF that tracks the benchmark while exposing the investor to a low (with large probability $p$ ) market risk level. While the mean-variance framework studies the trade-off between the absolute return and variance of a portfolio, our model analyzes the trade-off between absolute risk and relative return. Its objective function is defined in terms of the relative return (deviation from the return of the benchmark) of the EIF and is minimized subject to a probabilistic constraint on the overall variance of the EIF.

The motivation for the inclusion of the probabilistic constraint (5) is threefold. First, as aforementioned, we want to limit the consequences of the parameter estimation risk [4, 17, 19, 23, 24, 68, 70]. As indicated in [19, 29], the composition of the index fund is very sensitive to the values of the parameters, i.e., the asset returns and their VC matrix, and minor perturbations in their estimated values can lead to the construction of very different funds. The true values of the asset returns and of their variance-covariance matrix are unknown and unobservable, and multiple sources of errors (e.g., difficulty to obtain enough data points, instability of data, etc.) affect their estimation $[15,29]$. Despite this, most asset allocation models use a single-point estimate, such as the sample mean and the sample VC matrix of asset returns. This comes up to define the sample mean and sample VC matrix as deterministic parameters, thereby implicitly assuming that these are highly accurate estimates of their true counterparts. This high confidence in the accuracy of such estimates exacerbates the estimation risk. The need for developing asset allocation models that are less impacted by inaccuracies in the estimation of the moments of the asset returns has been recurrently stated (see, e.g., [19, 23, 29]). This gives us the impetus to model the asset returns and their covariance terms as random variables.

Second, the risk constraint (5) ensures that the return variability of the constructed EIF is low ( $\leq v$ ), with large probability $p$. Our approach can thus be viewed as a risk-averse form of enhanced indexation, since it tracks a market benchmark and limits the risk exposure. This is an important feature, since earlier studies showed that semi-active indexation strategies can result in funds that track very closely a benchmark but exhibit an overall variance larger than that of the benchmark [7, 29, 48, 49, 77]. Eighty-three percents of the EIFs analyzed by Jorion [48] have a larger risk than their benchmark. In order for an EIF and its benchmark to have similar risk profiles, several studies $[24,48]$ have suggested to minimize the tracking error while restricting the global variance of the fund. Note that our risk-averse enhancement differs from the previously proposed forms of enhanced indexation $[18,33,53]$ which track a benchmark as closely as possible while seeking an excess return.

Third, constraint (5) plays a critical role in controling the decisions of a fund manager and in avoiding agency problems reported by Jorion [49], in particular when the salary of fund managers includes a performance fee [40, 41, 49]. Performance fees are sometimes designed in such a way that, for a given return level of the managed fund, the manager's salary is higher if she takes more risk [41]. With the objective of limiting the moral hazard effect of performance-fee remuneration, Jorion proposes a portfolio optimization model that includes a
tracking error and a portfolio variance constraints [49]. He reports that the inclusion of the portfolio variance constraint significantly improves the performance of the allocation strategy. Our risk constraint (5) can be viewed as a covenant that prevents the construction of high-risk EIFs and limits the agency problem. In contrast to ours, Jorion's model is intended for the construction of a standard portfolio (i.e., no cardinality constraint) and is deterministic (single-point estimate of asset returns).

### 2.2 Risk Representation

In this section, we shall model the risk of the index fund and the covariance structure between asset returns. Three approaches have been widely used to derive the asset return VC matrix. The first one uses historical timeseries of asset returns to construct a sample VC matrix [77, 91]. For an asset universe of $n(n>0)$ securities, this approach involves the estimation of $n(n+1) / 2$ covariance terms and can lead to model specification issues, such as the obtaining of a VC matrix that is not positive semi-definite [29, 43]. The estimation of the VC matrix can also be affected by firm-specific events which momentarily influence several stocks but are unlikely to have a lasting effect on the behavior of asset returns [20]. The second method, called scenario method, consists in the generation of a finite set of scenarios representative of the future and in the estimation of the asset return vector and the VC matrix [31, 45, 72, 74, 91]. It allows for the introduction of expert views about the future return levels [67]. Such views may or may not be available and can be quite subjective. The third method involves the construction of a factor risk model $[16,20,52]$. It is based on the identification of a relatively small number of sources of risk, called factors, and on the quantification of the sensitivity of each asset return to each factor [71]. Factor models typically assume that asset returns depend linearly on the movement of a set of common factors and on the asset-specific return term and include an error term [45, 52, 67]. The risk induced by the volatility of the asset returns is decomposed into three elements: the systematic risk associated with the return of the common factors, the idiosyncratic risk specific to each asset, and the residual risk. The limited number of factors (see the three-factor Fama-French model [35] and [20] for a review) keeps the number of estimated factor covariance terms small, and allows for a more compact risk representation. Chan et al. [20] contend that factor models reduce the impact of the idiosyncratic risk in the VC matrix by relying on pervasive factors common to most assets. Three main families of factors (statistical, macroeconomic, and fundamental) exist. The reader is referred to [26, 27, 71] for a review of risk factor models and to [20] for a evaluation of the strengths and weaknesses of the possible methods for constructing a VC matrix.

Based on the above discussion, we propose a new stochastic factor risk model. We assume that the asset returns are driven by a factor model and that the returns of the factors are random variables. Denoting by $n$ and $m$ the respective number of risky assets and factors, the vector of random asset returns $\xi$ reads:

$$
\begin{equation*}
\xi=\mu+\beta^{\prime} f+\varepsilon, \tag{8}
\end{equation*}
$$

where $\mu \in \mathcal{R}^{n}$ is the vector of mean asset returns, $f \sim N(0, Q) \in \mathcal{R}^{m}$ represents the random factor returns, $\beta \in \mathcal{R}^{m \times n}$ is the factor loading matrix of the $n$ assets and $\varepsilon \sim N(0, D) \in \mathcal{R}^{n}$ is the residual return. The notation $y \sim N(a, O)$ denotes a normally distributed variable $y$ with mean $a$ and VC matrix $O$. Each component $\beta_{j i}$ of the factor loading matrix $\beta$ is the effect of factor $j$ on the return of asset $i$. Additional standard assumptions are that the factor returns are uncorrelated with the residual returns $\varepsilon$ and that the residual returns are mutually uncorrelated [20, 39, 45, 67]: $D$ is diagonal, positive semi-definite and its non-zero components are the residual return variances. Thus, the factor risk model implies that $\xi \sim N\left(\mu, \beta^{\prime} Q \beta+D\right)$, with $\beta^{\prime} Q \beta+D$ representing the VC matrix $\Sigma$ of the asset returns.

To derive an estimate of the vector $\mu$ and matrix $\beta$, we use time-series of observed asset returns $r_{i}^{t}$ and
observed factor returns $\tau_{j}^{t}$ for $l$ periods, with $l \gg m$, and we use standard linear regression (see also [39, 61]). The factor model (8) implies that:

$$
\begin{equation*}
r_{i}^{t}=\mu_{i}+\sum_{j=1}^{m} \beta_{j i} \tau_{j}^{t}+\varepsilon_{i}^{t}, \quad i=1, \ldots, n, t=1, \ldots l \tag{9}
\end{equation*}
$$

where each $\varepsilon_{i}^{t}, i=1, \ldots, n, t=1, \ldots, l$ is an independent normal random variable $N\left(0, \sigma_{i}^{2}\right)$.
Let $C=\left[\tau^{1} \cdots \tau^{l}\right] \in \mathcal{R}^{m}$ be the matrix of factor returns, $\varepsilon_{i}^{\prime}=\left[\varepsilon_{i}^{1}, \ldots, \varepsilon_{i}^{l}\right]$ be the vector of residual asset returns for $i, A=\left[e C^{\prime}\right]$, and $x_{i}^{\prime}=\left[\mu_{i}, \beta_{1 i}, \ldots, \beta_{m i}\right]$; we have that

$$
\begin{equation*}
y_{i}=\left(r_{i}^{1}, \ldots, r_{i}^{l}\right)^{\prime}=A x_{i}+\varepsilon_{i} . \tag{10}
\end{equation*}
$$

With $A$ having full column rank $\left(l \gg m\right.$ ), we can derive the least-square estimate $\hat{x}_{i}=\left(A^{\prime} A\right)^{-1} A^{\prime} y_{i}$ of $x_{i}$.

### 2.3 Reformulation of Stochastic Risk Constraint

Problem (SIP) is a stochastic integer programming problem. It belongs to the family of mixed-integer nonlinear programming problems (MINLP) whose continuous relaxation is non-convex. Besides the quadratic form of its objective function and its integrality requirements, the main computational challenge stems from the handling of the probabilistic constraint with random technology matrix (5). In its current form, problem (SIP) cannot be handled by any optimization solvers. Thus, in this section, we shall direct our efforts to the reformulation of (SIP) in a form that is amenable to its numerical solution.

### 2.3.1 Second-Order Cone Reformulation of Stochastic Constraint

Within the financial optimization literature, different approaches have been proposed for the reformulation of stochastic and robust constraints [11, 15, 39, 56, 61]. In [39], the authors propose a robust factor model to represent risk (see also [61]). They study various sources and forms of uncertainty, derive estimates of the underlying random variables, and construct confidence regions around the estimates. In [11, 15, 56], the risk representation is based on historical data of asset returns. Bodnar and Schmid [11] assume that the asset returns follow a normal distribution. In [15, 56], the only assumption on the asset returns is that their first and second moments are finite. In this case, a deterministic approximation can be derived using Cantelli's inequality. If the symmetry [15] or unimodality [56] of the distribution can be assumed, then tighter deterministic approximations modelled as second-order cone problems can be obtained using the Chebychev and Camp-Meidell probability inequalities.

The model proposed in this study does not assume anything regarding the probability distribution of the asset returns, except that they depend on the factor returns. In view of the difficulty to estimate the factor returns as well as their first and second moments, we model the estimated factor returns VC matrix $\hat{Q}$ as a stochastic one. The challenge of estimating the return VC matrix is acknowledged in the robust optimization literature [39, 42].

The probabilistic constraint (5) involves the computation of the variance of the index fund return that is unknown but that can be estimated with the factor model presented in Section 2.2. Since the VC matrix $\Sigma$ of asset returns is $\beta^{\prime} Q \beta+D$ (see Section 2.2), the estimated variance of the index fund return is

$$
\begin{equation*}
w^{\prime} \hat{\Sigma} w=w^{\prime}\left(\beta^{\prime} \hat{Q} \beta+\hat{D}\right) w, \tag{11}
\end{equation*}
$$

where $\hat{\Sigma}, \hat{Q}$, and $\hat{D}$ denote the estimated VC matrices of asset, factor and residual returns, respectively.
We shall now analyze the distributional properties of the estimated VC matrices $\hat{Q}$ and $\hat{D}$. This will enable us to determine the probability distribution of the estimated VC matrix $\hat{\Sigma}$ (Proposition 1) and of $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$
(Theorem 1). The knowledge of the distribution of $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$ will lead to Proposition 2 in which we will derive a deterministic formulation equivalent to the stochastic constraint (5).

The estimate of the matrix of factor returns is obtained by using a series of historical factor returns. As indicated in Section 2.2 and similarly to [39, 61, 78, 79], the vector $f^{t}$ of factor returns at time $t$ is normally distributed with mean 0 and VC matrix $Q$, and the vectors of factor returns at different periods are independent of each other. There is no independence assumption between the returns of factors $j$ and $j^{\prime}$ at the same period $t$ ( $f_{j}^{t}$ and $f_{j^{\prime}}^{t}$ are not independent). The notation $W_{a}(b, C)$ refers to the $a$-dimensional Wishart distribution with $b$ degrees of freedom and variance-covariance matrix $C$.

Proposition 1 If $\eta_{1}, \ldots, \eta_{l}$ are independent $n$-variate normal random vectors $N(\mu, \Sigma)$ and $l>n>0$, the sample covariance matrix $\hat{\Sigma}$ has a Wishart density function $W_{n}\left(l-1, \frac{1}{l-1} \Sigma\right)$ [73].
Thus, the density function of $(l-1) \hat{\Sigma}$ is $W_{n}(l-1, \Sigma)$. Proposition 1 implies that the estimated VC matrices $\hat{Q}$ and $\hat{D}$ of normally distributed factor $f$ and residual $\varepsilon$ returns follows the Wishart distributions $W_{m}\left(l-1, \frac{1}{l-1} Q\right)$ and $W_{n}\left(l-1, \frac{1}{l-1} D\right)$. Theorem 1 defines the distribution of $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$ and is based on Proposition 1.

Theorem 1 If $\eta_{1}, \ldots, \eta_{l}$ are independent random variables following an identical $n$-variate normal distribution $N(\mu, \Sigma)$ with an estimated VC matrix $\hat{\Sigma}$, then $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$ has a $\chi^{2}$ distribution with $(l-1)$ degrees of freedom.

Proof: The following result was shown in [73]: If the $[n \times n]$-dimensional random matrix $A$ follows $W_{n}(m, \Sigma)$ and $Y$ is an $[n \times q]$-dimensional matrix, then $Y^{\prime} A Y$ follows the Wishart distribution $W_{q}\left(m, Y^{\prime} \Sigma Y\right)$ and $\frac{Y^{\prime} A Y}{Y^{\prime} \Sigma Y}$ follows a Wishart distribution $W_{q}(m, 1)$.
Thus, with $Y=w$ and $A=(l-1) \hat{\Sigma}$, the distribution of $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$ is $W_{1}(l-1,1)$, which is the $\chi^{2}$ distribution with $(l-1)$ degrees of freedom.
Theorem 1 implies that the distribution of $\beta^{\prime} \hat{Q} \beta$ is $W_{n}\left((l-1), \frac{1}{l-1} \beta^{\prime} Q \beta\right)$. As the sum of independent random matrices with Wishart distributions is a matrix following a Wishart distribution [73], the probability distribution of $\beta^{\prime} \hat{Q} \beta+\hat{D}$ is $W_{n}\left((l-1), \frac{1}{l-1}\left(\beta^{\prime} Q \beta+D\right)\right)$. Thus, it results from Theorem 1 that $w^{\prime} \hat{\Sigma} w$ in (11) follows the Wishart distribution $W_{1}\left((l-1), \frac{1}{l-1} w^{\prime}\left(\beta^{\prime} Q \beta+D\right) w\right)$ and that $\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$ is $W_{1}(l-1,1)$, which is the $\chi^{2}$ distribution with $(l-1)$ degrees of freedom. This leads to Proposition 2 which defines a deterministic formulation equivalent to the probabilistic constraint.

Proposition 2 Denoting by $F_{\zeta}^{-1}(1-p)$ the $(1-p)$-quantile of the probability distribution $F_{\zeta}$ of $\zeta=\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w}$,

$$
\begin{equation*}
\frac{v F_{\zeta}^{-1}(1-p)}{l-1}-w^{\prime} \hat{\Sigma} w \geq 0 \tag{12}
\end{equation*}
$$

is a deterministic constraint and is equivalent to the stochastic constraint $\mathbb{P}\left(w^{\prime} \Sigma w \leq v\right) \geq p$.
Proof: Since $w^{\prime} \Sigma w \neq 0$ and $v \neq 0$, the left-hand side of the stochastic inequality (5) can be rewritten as:

$$
\begin{equation*}
\mathbb{P}\left(w^{\prime} \Sigma w \leq v\right)=\mathbb{P}\left(\frac{1}{w^{\prime} \Sigma w} \geq \frac{1}{v}\right)=\mathbb{P}\left(\frac{(l-1) w^{\prime} \hat{\Sigma} w}{w^{\prime} \Sigma w} \geq \frac{(l-1) w^{\prime} \hat{\Sigma} w}{v}\right)=1-\mathbb{P}\left(\zeta \leq \frac{(l-1) w^{\prime} \hat{\Sigma} w}{v}\right) \tag{13}
\end{equation*}
$$

Denoting by $F_{\zeta}^{-1}$ the inverse of the probability distribution $F_{\zeta}$ and using the relationship (13), (5) becomes:

$$
\begin{align*}
1-\mathbb{P}\left(\zeta \leq \frac{(l-1) w^{\prime} \hat{\Sigma} w}{v}\right) \geq p & \Leftrightarrow 1-p \geq F_{\zeta}\left(\frac{(l-1) w^{\prime} \hat{\Sigma} w}{v}\right) \quad \Leftrightarrow F_{\zeta}^{-1}(1-p) \geq \frac{(l-1) w^{\prime} \hat{\Sigma} w}{v}  \tag{14}\\
& \Leftrightarrow \frac{v F_{\zeta}^{-1}(1-p)}{l-1}-w^{\prime} \hat{\Sigma} w \geq 0 \tag{15}
\end{align*}
$$

Constraint (15) is equivalent to (5) and does not include any random number. Thus, (15) is a deterministic constraint enforcing the same requirements as the stochastic constraint (5).

The deterministic equivalent problem ( $\mathbf{( 1 1 )}$ of the stochastic problem (SIP) is thus:

$$
\begin{array}{rc}
\quad(\mathbf{P 1}): & \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \\
\text { subject to } & \frac{v F_{\zeta}^{-1}(1-p)}{l-1}-w^{\prime} \hat{\Sigma} w \geq 0  \tag{17}\\
& (2)-(4),(6),(7) .
\end{array}
$$

### 2.3.2 Properties

In order to evaluate the computational tractability of problem (P1), we shall now analyze the convexity properties of the feasible set defined by constraint (12) and by the continuous relaxation of problem (P1).

Proposition 3 Constraint (12) defines a convex feasible set and is a second-order cone constraint.
Proof: To show that (12) is a second-order cone constraint, thus defining a convex feasible set, we must demonstrate that the left-hand side of (12) is concave in $w$. Since $\frac{v F_{\zeta}^{-1}(1-p)}{l-1}$ is a constant ( $p$ is fixed), it comes up to showing that $w^{\prime} \hat{\Sigma} w$ is convex.
From (11), we know that $w^{\prime} \hat{\Sigma} w=w^{\prime} \beta \hat{Q} \beta^{\prime} w+w^{\prime} \hat{D} w$. The first term is convex, since the VC matrix $\hat{Q}$ of the factor return is positive semidefinite. The second term $w^{\prime} \hat{D} w=\sum_{i=1}^{N} w_{i}^{2} d_{i i}$ is convex, since $\hat{D}$ is a diagonal matrix and its diagonal terms $d_{i i}$ are non-negative numbers representing the variance of the residual returns. Provided that the convexity property carries over from terms to sum, we have the result that was set out to prove.

Proposition 3 implies that:
Corollary 1 The continuous relaxation of the deterministic equivalent problem (P1) is a convex problem.
Each of the linear and second-order cone constraints in the continuous relaxation of (P1) defines a convex feasible region. The intersection of convex sets is convex. The objective function is the sum of the squared return differences and is thus quadratic and convex. The minimization of a convex function over a convex feasible set is a convex programming problem.

### 2.4 Block Decomposition of Variance-Covariance Matrix

The estimated VC matrix of asset returns $\beta^{\prime} \hat{Q} \beta+\hat{D}$ is a very dense $[n \times n]$-dimensional matrix. The risk constraint (12) involves the calculation of the estimated variance of the index fund return (11). This requires the computation of $(n(n+1) / 2)$ second-degree terms, which is a very computationally intensive task when the asset universe includes a large number $n$ of securities. To ease the computations, we reformulate the VC matrix through the introduction of $m$ (one per factor $j$ ) fictitious assets [45, 83]. Each fictitious position $w_{j}^{L}$ is constructed as a linear combination of the positions in the $n$ risky securities

$$
\begin{equation*}
w_{j}^{L}=\sum_{i=1}^{n} \beta_{j i} w_{i}, j=1, \ldots, m \tag{18}
\end{equation*}
$$

and can be interpreted as the average responsiveness of the index fund to the factor $j$. Indeed, the coefficient $\beta_{j i}$ multiplying a position $w_{i}$ is the sensitivity of asset $i$ to factor $j$. The resulting augmented model gives a new VC matrix allowing for a much faster computation of the risk of the index fund.

Combining the "real" and fictitious asset positions, we obtain the augmented $(n+m)$-dimensional vector $\mathbf{w}$ such that $\mathbf{w}^{\prime}=\left[w, w^{L}\right]$. The estimated variance of the index fund return becomes

$$
\begin{equation*}
w^{\prime} \hat{\Sigma} w=\mathbf{w}^{\prime} \hat{V}^{L} \mathbf{w} \tag{19}
\end{equation*}
$$

where the augmented VC matrix $\hat{V}^{L}=\left(\begin{array}{cc}\hat{D} & 0 \\ 0 & \hat{Q}\end{array}\right)$ is almost diagonal and much sparser than in its original form $\beta^{\prime} \hat{Q} \beta+\hat{D}$. Using $\hat{V}^{L}$, the computation of the variance of the index fund return (19) requires the calculation of $\left(n+\frac{m(m+1)}{2}\right)$ second-degree terms instead of $\frac{n(n+1)}{2}$ terms when $\beta^{\prime} \hat{Q} \beta+\hat{D}$ is used as VC matrix. This alleviates the computational burden, since $\left(n+\frac{m(m+1)}{2}\right) \geq \frac{n(n+1)}{2}$ for $n \geq m+1$, and $n$ is typically much larger than $m$.

The transformation described above and based on the introduction of $m$ fictitious assets provides a second deterministic equivalent formulation (P2) for (SIP):

$$
\begin{array}{cc}
\text { (P2) : } & \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \\
\text { subject to } & \mathbf{w}^{\prime} \hat{V}^{L} \mathbf{w} \leq \frac{v F_{\zeta}^{-1}(1-p)}{l-1}  \tag{21}\\
& (2)-(4),(6),(7) .
\end{array}
$$

The fictitious positions $w^{L}$ are unrestricted in sign. Their introduction does not affect the features of the deterministic equivalent problem. The constraint (21) is also a second-order cone constraint, and the continuous relaxation of problem ( $\mathbf{P} 2$ ) is a convex programming problem. Section 4.2.1 reports the time needed to solve the continuous relaxations of the two deterministic equivalent problems ( $\mathbf{P 1}$ ) and ( $\mathbf{P} 2$ ), and discusses the computational advantage of using the VC matrix obtained with the block-decomposition method.

### 2.5 Model with Buy-in Threshold Constraints

In this section, we introduce buy-in threshold constraints (23) which prevent the holding of small positions in the index fund. Investors are reluctant to hold small positions, since they have poor liquidity and have little impact on the return of the index fund, but generate brokerage fees and monitoring costs [47]. Buy-in threshold constraints set a lower bound $w_{\text {min }}$ on the proportion of capital invested in any asset included in the index fund and remove the negative effect of small positions. Along with (3), (23) force each $\tilde{w}_{i}$ to take value 0 or a value at least equal to $w_{\text {min }}$. The corresponding deterministic equivalent problem $(\mathbf{P B})$ reads:

$$
\begin{array}{rc}
\quad(\mathbf{P B}): & \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \\
\text { subject to } & (2)-(4),(6),(7),(21)  \tag{23}\\
& w_{\min } \gamma \leq \tilde{w} .
\end{array}
$$

As compared to ( $\mathbf{P} 2$ ), (PB) contains $(n+1)$ additional buy-in threshold constraints with the binary variables $\gamma$.

## 3 Solution Method

The reformulation of the stochastic constraint provides a deterministic mixed-integer nonlinear programming problem (MINLP) whose continuous relaxation is convex. Algorithmic techniques used to solve MINLP problems are most often based on relaxation schemes. Solution approaches based on nonlinear branch-and-bound
algorithms (see, e.g., $[13,15,75]$ ) or on outer approximations (see, e.g., $[34,36,57,89]$ ) have been proposed. A recent study [14] provides a comprehensive review of the recently released MINLP solvers (Bonmin [13], Couenne [6], FilMINT [1]) and the families of algorithmic techniques used to solve MINLP problems [36, 75].

In Sections 3.1 and 3.2, we propose two variants of an exact outer approximation method which converges to the optimal solution in a finite number of iterations. The computational challenges raised by problems ( $\mathbf{P 1}$ ) and (P2) stem from the quadratic objective function, the second-order cone constraint, and the integrality restrictions linked with the cardinality constraint. Many outer approximation algorithms derive a polyhedral representation of the lower envelope of the non-linear constraint [34, 57, 89]. The continuous relaxation of problem (P2) is a second-order cone problem which can be very efficiently solved with interior point algorithms (see Section 4.2.1). That is why, instead of constructing a polyhedral relaxation of the nonlinear constraints, our algorithmic approach focuses on the combinatorial challenges posed by the cardinality constraint.

The proposed outer approximation approach is based on a hierarchical organization of the computations with expanding sets of binary-restricted variables. The general idea resides in the construction of a simpler approximation problem whose feasible set includes the one of the original problem. The outer approximation problem is obtained by relaxing the integrality conditions on a subset of the binary variables and by reformulating the cardinality constraint (4). At each iteration $t$, we partition the set of binary variables into two subsets $S_{1}^{(t)}$ and $S_{2}^{(t)}$, and remove the integrality restrictions on the variables included in $S_{2}^{(t)}$. The partitioning is a critical step in the algorithm and is designed in such a way that the set of variables for which the integrality restrictions are maintained is $(i)$ small enough to make the approximation problem easy to solve, and, at the same time, is (ii) large enough to contain, with a high probability, the securities included in the optimal index fund. The solution of the approximation problem provides a lower bound (i.e., minimization problem) on the optimal solution of the original problem and is used to define the stopping criterion. If the optimal solution of the approximation problem cannot be proven feasible, and thus optimal, for the original problem, the approximation problem is tightened. This is done by enlarging the set $S_{1}^{(t)}$ of variables on which the binary restrictions are restored: $S_{1}^{(t)} \subseteq S_{1}^{(t+1)}$. This results in a series of increasingly tighter outer approximations and guarantees the finite convergence of the algorithm. The next sub-sections provide a more elaborate description of the algorithm and show that it is simple to implement and general enough to be applied to other types of problems.

### 3.1 Outer Approximation Algorithm $O A A$

The exact outer approximation algorithm OAA involves an initialization (Section 3.1.1) and an iterative (Section 3.1.2) phases. Each iteration terminates with the verification of whether the stopping criterion is met (37). The properties and the pseudo-code of the algorithm are presented in Section 3.1.3.

### 3.1.1 Initialization

The successive outer approximation problems are obtained by relaxing the integrality requirements on a subset of the binary variables $\gamma_{i}$. The initial outer approximation problem $(\mathbf{O A})^{(0)}$ is the nonlinear continuous relaxation

$$
\begin{array}{cc}
(\mathbf{O A})^{(0)}: & \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \\
\text { subject to } & (2)-(4),(6),(21) \\
& 0 \leq \gamma \leq 1 \tag{25}
\end{array}
$$

of the deterministic equivalent problem ( $\mathbf{P 2}$ ) (Instead of (P2), we could also use (P1)). The notation $\gamma^{*(t)}$ is used thereafter to refer to the value taken by $\gamma$ in the optimal solution of the outer approximation problem $(\mathbf{O A})^{(t)}$. The same notation style will be used for all decision variables ( $\tilde{w}$ and $\gamma_{S}$ ).

The optimal solution of $(\mathbf{O A})^{(0)}$ is used to initiate the sequence of outer approximations. If all the variables $\gamma_{i}, i=1, \ldots, n$ take an integer value in the optimal solution $\left(\tilde{w}^{*(0)}, \gamma^{*(0)}\right)$ of $(\mathbf{O A})^{(0)}$ in which the integrality restrictions on $\gamma(25)$ are relaxed, then $\left(\tilde{w}^{*(0)}, \gamma^{*(0)}\right)$ is also optimal for (P2), and we stop. Otherwise, we start the iterative process and partition of the set of securities into:

$$
\begin{align*}
S_{1}^{(0)} & =\left\{i: \tilde{w}_{i}^{*(0)}>0, i=1, \ldots, n\right\}  \tag{26}\\
S_{2}^{(0)} & =\left\{i: \tilde{w}_{i}^{*(0)}=0, i=1, \ldots, n\right\} . \tag{27}
\end{align*}
$$

The integrality restrictions on the binary variables $\gamma_{i}, i \in S_{1}^{(0)}$ are maintained, while they are removed on those in $S_{2}^{(0)}$. The idea is to maintain the integrality restrictions on the variables associated with the assets in which the investor would hold positions if there was no cardinality constraint (4). The approach is based on our preliminary experiments. We learnt from these that a security $i$ with variable $w_{i}$ taking a positive (resp., null) value in the optimal solution of the relaxed problem has a good chance (resp., is unlikely) to be included in the optimal EIF. In other words, the optimal solution of $(\mathbf{O A})^{(0)}$ is very "informative" as for the composition of the optimal EIF. Similar findings were found in previous asset allocation studies subject to a cardinality constraint [51, 65].

### 3.1.2 Iterative Process

At each iteration $t$, the optimal solution of the outer approximation problem $(\mathbf{O A})^{(t-1)}$ is used to update the composition of the sets $S_{1}^{(t)}$ and $S_{2}^{(t)}$ :

$$
\begin{align*}
S_{1}^{(t)} & =S_{1}^{(t-1)} \bigcup\left\{i: \tilde{w}_{i}^{*(t-1)}>0, i \in S_{2}^{(t-1)}\right\}  \tag{28}\\
S_{2}^{(t)} & =\{1,2, \ldots, n\} \backslash S_{1}^{(t)} \tag{29}
\end{align*}
$$

The notation $\tilde{w}^{*(t-1)}$ is the vector of optimal positions in problem $(\mathbf{O A})^{(t-1)}$ and $S_{1}^{(0)}$ and $S_{2}^{(0)}$ are defined by (26) and (27), respectively. The set updating process (28)-(29) serves as basis for the formulation of the current outer approximation problem $(\mathbf{O A})^{(t)}$. Problem $(\mathbf{O A})^{(t)}$ is such that: $(i)$ a binary variable (34) is associated with each asset included in $S_{1}^{(t)}(31)$; (ii) there is no binary variable individually assigned to any of the assets included in the set $S_{2}^{(t)}$. Instead, we associate a binary variable $\gamma_{S}$ (35) with the group of assets included in $S_{2}^{(t)}$ (32): $\gamma_{S}$ takes value 1 if any one of the variables $\tilde{w}_{i}, i \in S_{2}^{(t)}$ is strictly positive.
$(\mathbf{O A})^{(t)}: \min \left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right)$
subject to
(2), (6), (21)

$$
\begin{array}{cc}
\tilde{w}_{i} \leq \gamma_{i}, & i \in S_{1}^{(t)} \\
\sum_{i \in S_{2}^{(t)}} \tilde{w}_{i} \leq \gamma_{S} & \\
\sum_{i \in S_{1}^{(t)}} \gamma_{i}+\gamma_{S} \leq K & \\
\gamma_{i} \in\{0,1\}, & i \in S_{1}^{(t)} \\
\gamma_{S} \in\{0,1\} . & \tag{35}
\end{array}
$$

The approximation problem $(\mathbf{O A})^{(t)}$ contains $\left|S_{1}^{(t)}\right|+1$ binary variables, while the original problem contains $n$ binary variables. In the next $(t+1)$ iteration, the assets $i$ for which $w_{i}^{(t) *}>0$ are moved to $S_{1}^{(t+1)}(28)$ and a new binary variable is introduced for each of them.

### 3.1.3 Properties

Proposition 4 Problem $(\mathbf{O A})^{(t)}$ is an outer approximation of the deterministic problem $(\mathbf{P 2})$ equivalent to $(\mathbf{S I P})$.
Proof It is enough to show that the feasible set defined by

$$
\left\{\left(\tilde{w}, \gamma, \gamma_{S}\right):(31),(32),(33),(34),(35)\right\}
$$

is a relaxation of the feasible set defined by

$$
\begin{equation*}
\{(\tilde{w}, \gamma):(3),(4),(7)\} \tag{36}
\end{equation*}
$$

and does not cut any feasible solution for the set (36).
Constraint (32) forces $\gamma_{S}$ to take value 1 if an investment is made in any of the assets $i$ assigned to $S_{2}^{(t)}$. Thus, (33) is a relaxation of the cardinality constraint (4): (33) allows investing in at most $K$ (resp., $K-1$ ) of the assets included in $S_{1}^{(t)}$ if no (resp., at least one) position is held in any of the securities included in $S_{2}^{(t)}$, while (4) allows any investment in at most $K$ assets $i, i=1, \ldots, n$. Since $\{1,2, \ldots, n\}=S_{2}^{(t)} \cup S_{1}^{(t)}$ and $S_{2}^{(t)} \cap S_{1}^{(t)}=\emptyset$ (see (28), (29)), we obtain the result which was set out to prove.

Proposition 5 The optimal solution $\left(\tilde{w}^{*(t)}, \gamma^{*(t)}, \gamma_{S}^{*(t)}\right)$ of $(\mathbf{O A})^{(t)}$ is optimal for (P2) if

$$
\begin{equation*}
|Q| \leq K, \quad \text { with } \quad Q=\left\{i: \tilde{w}_{i}^{*(t)}>0, i=1, \ldots, n\right\} \tag{37}
\end{equation*}
$$

Proof Problem $(\mathbf{O A})^{(t)}$ relaxes some of the integrality restrictions and the cardinality constraint (4) imposed in (P2). Thus, if the optimal solution of $(\mathbf{O A})^{(t)}$ satisfies the cardinality constraint, which is the case if (37) holds, it is feasible and optimal for (P2), since the feasibility set of $(\mathbf{O A})^{(t)}$ includes that of (P2).
The above condition (37) is used as the stopping criterion for the algorithmic process.
Proposition 6 The algorithm OAA generates a series of increasingly tighter outer approximations.
Proof It is an immediate consequence of the updating process of the sets $S_{1}^{(t)}$ and $S_{2}^{(t)}$ using (28) and (29). At iteration $t$, all the variables $\gamma_{i}, i \in S_{1}^{(t-1)}$ defined as binary at iteration $(t-1)$ remain defined as binary, while at least one of $\gamma_{i}, i \in S_{2}^{(t-1)}$ on which the integrality restriction was relaxed at $(t-1)$ is defined as binary at $t$.

Proposition 7 The algorithm OAA has the finite convergence property. It terminates after at most $(n-K-1)$ iterations.

Proof The optimal solution $\left(\tilde{w}^{*(0)}, \gamma^{*(0)}\right)$ defines the set $S_{1}^{(0)}=\left\{i: \tilde{w}_{i}^{*(0)}>0, i=1, \ldots, n\right\}$. If $\left|S_{1}^{(0)}\right| \leq K$, then the solution is also optimal for (P2) and we stop. Otherwise, if $\left|S_{1}^{(0)}\right|>K$, then we move to the next iteration in which the integrality restrictions are restored on at least one of the $\gamma_{i}, i \in S_{2}^{(0)}$ that were so far defined as continuous. Thus, it is clear that the algorithmic process finds the optimal solution of ( $\mathbf{( P 2 )}$ in at most ( $n-K-1$ ) iterations.

Proposition 6 indicates that each new outer approximation problem comprises a larger number of binary variables and is likely more challenging to solve. For our algorithmic approach to be efficient, it is crucial that the number of iterations remains small, which underlines the importance of the selection of the variables for which integrality restrictions are enforced. This question is investigated in Section 4. The pseudo-code of the OAA algorithm follows.

## Outer Approximation Algorithm OAA

## Initialization:

$t:=0$;
Solve the continuous relaxation $(\mathbf{O A})^{(0)}$ of (P2). Let $\left.\left(\tilde{w}^{*(0)}, \gamma^{*(0)}\right):=\operatorname{argmin}(\mathbf{( O A})^{0}\right)$;
Define $S_{1}^{(0)}$ and $S_{2}^{(0)}$ according to (26) and (27);
if $\left|S_{1}^{(0)}\right| \leq K$ then
$\left(\tilde{w}^{*(0)}, \gamma^{*(0)}\right)$ is optimal for (P2)
end
else
Iterative Process:
repeat
$t:=t+1 ;$
Let $\left.\left(\tilde{w}^{*(t-1)}, \gamma^{*(t-1)}, \gamma_{S}^{*(t-1)}\right):=\operatorname{argmin}(\mathbf{( O A})^{t-1}\right) ;$
Update $S_{1}^{(t)}$ and $S_{2}^{(t)}$ as in (28) and (29), respectively ;
Solve (OA) ${ }^{(t)}$;
until $|Q| \leq K$ (37);
end

### 3.2 Outer Approximation Algorithm with Optimality Cut OAC

The algorithm OAC presented in this section differs from the algorithm OAA (Section 3.1) as follows: (i) it involves the construction of an inner approximation problem whose optimal solution provides an optimality cut, also called objective level cut [58, 59], for the original problem (P2); (ii) the optimality cut is introduced in the formulation of the successive outer approximation problems.

The optimality cut is derived as follows. First, we solve the continuous relaxation $\mathbf{O A}{ }^{(0)}$ of problem ( $\mathbf{P 2}$ ). Its optimal solution defines the composition of the sets $S_{1}^{(0)}$ and $S_{2}^{(0)}$ (see (26)-(27)). Second, we construct an index fund by only considering the assets $i$ included in $S_{1}^{(0)}$. This is achieved by solving problem (IA).

Proposition 8 The nonlinear optimization problem

$$
\begin{array}{ccc}
(\mathbf{I A}): & \min \left(r^{\prime} \bar{w}-r_{M}\right)^{\prime}\left(r^{\prime} \bar{w}-r_{M}\right) & \\
\text { subject to } & \sum_{i \in S_{1}^{(0)}} \bar{w}_{i}=1 & \\
\overline{\mathbf{w}}^{\prime} \hat{V}^{L} \overline{\mathbf{w}} \leq \frac{v F_{\zeta}^{-1}(1-p)}{l-1} &  \tag{40}\\
& \bar{w}_{i} \leq \gamma_{i}, & i \in S_{1}^{(0)} \\
\sum_{i \in S_{1}^{(0)}} \gamma_{i} \leq K & \\
& \bar{w}_{i} \geq 0, & i \in S_{1}^{(0)} \\
& \gamma_{i} \in\{0,1\}, & i \in S_{1}^{(0)} .
\end{array}
$$

is an inner approximation of problem (P2).
Proof Problem (IA) is subject to the same constraints as (P2), but must satisfy them by using only a subset $\left(i, i \in S_{1}^{(0)}\right)$ of the securities considered in (P2). Thus, the feasible set of (IA) is included in the one defined by
(P2): (IA) is an inner approximation of (P2).
Problem (IA) is much easier to solve than (P2), since (IA) has the same features as (P2) (i.e., its continuous relaxation is convex), but comprises a much smaller number of decision variables $\bar{w}_{i}$ and $\gamma_{i}, i \in S_{1}^{(0)}$. We denote by $\mathbf{u}_{\mathbf{I A}}^{*}$ the optimal value of (IA).

## Proposition 9 The inequality

$$
\begin{equation*}
\left(r^{\prime} \tilde{w}-r_{M}\right)^{\prime}\left(r^{\prime} \tilde{w}-r_{M}\right) \leq \mathbf{u}_{\mathbf{I} \mathbf{A}}^{*} \tag{45}
\end{equation*}
$$

is a objective level cut for (P2).
Proof Proposition 8 establishes that (IA) is an inner approximation of (P2). Any feasible solution for (IA) is feasible for ( $\mathbf{P} 2$ ) and the optimal value $\mathbf{u}_{\mathbf{I A}}^{*}$ of (IA) is an upper bound (minimization problem) on the optimal value of ( $\mathbf{P} 2$ ). Therefore, (45) does not cut any solution that can be optimal for (P2).
The next step is to introduce the cut (45) in the successive outer approximation problems $\mathbf{O A}{ }^{(t)}$. The cut (45) allows the elimination of feasible, yet non-optimal solutions for $(\mathbf{P} 2)$. Clearly, the objective pursued here is to be able to quickly generate a cut of "reasonable" quality. The cut is expected to help prune the branch-and-bound tree and speed up the finding of the optimal solution of problem $\mathbf{( P 2 )}$. Section 4 evaluates the contribution of the incorporation of the cut into the formulation of the outer approximation problems.

## 4 Computational Results

### 4.1 Test Problems

The market index that we attempt to replicate is the Standard \& Poor's (S\&P) 500 index. As factors, we use the Fama-French [35] factors (excess return on the market, small-minus-big return, high-minus-low return, momentum of the market), the risk free interest rate, and two major market indices (Nasdaq and Dow Jones Composite Average). The choice of the factors is based on the existing literature [16, 21, 28, 35, 39, 61, 74] in order to obtain a model with good predictability. The selection of the "optimal" set of factors is beyond the scope of this paper.

The monthly returns of more than 2000 stocks are extracted from the Compustat database (from January 1997 to December 2006). The Fama-French factors and the risk free interest rate come from the Kenneth French Data Library, while the return time-series of the S\&P 500, Nasdaq and Dow Jones Composite Average indices are obtained from the Center for Research in Security Prices. We partition the data points into a training and testing set. The training (in-sample) period ranges from January 1997 to December 2005. The training set includes the 108 corresponding monthly return data points. The testing (out-of-sample) period is from January 2006 to December 2006 and the testing set contains 12 monthly return data points. All the EIFs presented in Section 4.2 that focuses on the computational performance of the proposed algorithms are constructed by using the training data only. The out-of-sample data are used in Section 4.3 for cross-validation purposes and to analyze the performance of the constructed funds on the out-of-sample period.

We build five families of problem instances (Table 1) which differ in the size $n$ (up to 1000 assets) of the market universe and the number $K$ of assets that can be included in the index fund. Eight data sets were generated for each instance family. We consider several reasonable values of $K(0.03 n \leq K \leq 0.05 n)$ to verify whether the value of $K$ affects the performance of the proposed algorithms. For each problem instance, the securities included in the asset universe are selected randomly and $p$ is set to $90 \%$. Each problem is modelled with the AMPL modeling language and solved on a 64-bit Dell Precision T5400 Workstation with Quad Core Xeon Processor X5460 3.16GHz CPU and 4X2GB of RAM. Each problem instance is solved with six algorithmic methods:

- the outer approximation method without optimality cut (Section 3.1) used in conjunction with the standard branch-and-bound (B\&B) algorithm of the Bonmin solver, the one with optimality cut (Section 3.2) used with the standard $B \& B$ algorithm of the Bonmin solver, and the standalone nonlinear $B \& B$ algorithm of the Bonmin solver. The three solution approaches are thereafter referred to as $O A A-B, O A C-B$ and Bonmin, respectively. Among the algorithms implemented within Bonmin, preliminary tests showed that the $\mathrm{B} \& \mathrm{~B}$ algorithm is the most efficient one for this type of problems. The nonlinear continuous relaxations were solved with the interior point solver ipopt;
- the outer approximation method without optimality cut (Section 3.1) used in conjunction with the standard B\&B algorithm of the Cplex 12.1 solver, the one with optimality cut (Section 3.2) used with the standard $\mathrm{B} \& \mathrm{~B}$ algorithm of the Cplex 12.1 solver, and the standalone nonlinear $\mathrm{B} \& \mathrm{~B}$ of the Cplex 12.1 solver. The solution approaches are thereafter referred to as $O A A-C, O A C-C$ and $C p l e x$, respectively.

Table 1: Families of Instances

| $n$ | 100 | 200 | 500 | 700 | 1000 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $K$ | 5 | 10 | 25 | 25 | 30 |

### 4.2 Algorithmic Evaluation

This section is decomposed into three main sub-sections. The first sub-section compares the solution times for the continuous relaxations of the two deterministic equivalent problems (P1) and (P2). The second sub-section evaluates the computational efficiency, the robustness, and the scalability of the proposed outer approximation method. The third sub-section analyzes the model with buy-in threshold constraints.

### 4.2.1 Impact of Block Decomposition of Variance-Covariance Matrix

In this section, we evaluate the contribution of the block decomposition method by comparing the computational times for the solution of the continuous relaxations (RP1) and (RP2) of the deterministic equivalent problems $(\mathbf{P 1})$ and (P2). Recall that the only difference between (P1) and (P2) resides in the formulation of the variancecovariance matrix and in the fact that formulation ( $\mathbf{P} 2$ ) is based on the proposed block decomposition method.

For the 40 instances, the optimal solution of (RP2) is obtained (much) faster than that of (RP1). This conclusion applies with both the Bonmin and Cplex 12.1 solvers. For each family of problem instances, we display the average relative time gain obtained by solving (RP2) instead of (RP1) with Bonmin (Figure 1) and with Cplex 12.1 (Figure 2). This highlights the tremendous contribution of the block decomposition method. With Bonmin, the average time to solve the (RP2) formulation for the eight 100 -asset instances is 0.14453 seconds, while it is equal to 0.19141 seconds with the (RP1) formulation. This represents an average relative gain of $24.49 \%$. Most importantly, the relative time gain significantly increases as the size and the complexity of the problem increase, culminating to $85.90 \%$ for the 1000 -asset instances solved with Cplex 12.1 . With Cplex 12.1, the average solution time for the (RP2) formulation of the 1000 -asset instances is 5.56 seconds, whereas it is 43.40 seconds with the (RP1) formulation. Given that the optimal solution of the integer problems ( $\mathbf{P 1}$ ) and ( $\mathbf{P 2}$ ) requires the solution of multiple (one at each node in the branch-and-bound tree) continuous problems, the critical impact of the block decomposition method is evident.

### 4.2.2 Comparison of Solution Methods

In this section, we compare the computational efficiency of the six solution methods described in Section 4. The details of the computational results are provided in the Appendix. Tables 4 and 5 report the optimal objective


Figure 1: Average Relative Gain in Computational Time with the Block Decomposition Method - Bonmin


Figure 2: Average Relative Gain in Computational Time with the Block Decomposition Method - Cplex 12.1
value, the number of iterations needed with $O A A-B, O A C-B, O A A-C$ and $O A C-C$, the integrality and optimality gaps, the best solution found, and the times needed to find the optimal solution and to prove that it is optimal. The main objective of this section is to verify whether the solution process for the construction of EIFs is improved by using the proposed outer approximation framework as opposed to using the Bonmin and Cplex 12.1 solvers on a standalone basis. In what follows, excerpts from Table 4 and 5 are used to answer this question and shed light on the quality of the solutions, the computational efficiency and the scalability of the six algorithms.

Solution quality: Figure 3 shows the number of instances for which the optimal solution is found within one hour with the three methods used with the Bonmin solver. Figure 4 provides the same information for the three algorithms used with the Cplex 12.1 solver. Figure 5 displays the number of instances for which the optimality of the best solution found is proven within one hour with the three algorithms used with the Bonmin solver. Figure 6 provides the same information for the three algorithms used with the Cplex 12.1 solver.


Figure 3: Solution Quality with Bonmin: Percentage of Instances when the Optimal Solution is Found (one hour)


Figure 5: Solution Quality with Cplex 12.1: Percentage of Instances when the Optimal Solution is Found (one hour)


Figure 4: Solution Quality with Bonmin: Percentage of Instances when Optimality is Proven (one hour)


Figure 6: Solution Quality with Cplex 12.1: Percentage of Instances when Optimality is Proven (one hour)

The outer approximation algorithms $O A A-B$ and $O A C-B$ find the optimal solution and prove optimality for all the 100 -asset and 200-asset instances. On the other hand, Bonmin does not find the optimal solution for
$12.5 \%$ of the 200 -asset instances and is not able to prove that the solution is optimal for $12.5 \%$ of the 100 asset and 200-asset instances. The Cplex 12.1 solver, regardless of whether it is used on a standalone basis (Cplex) or complemented with the OAA or OAC approaches, finds the optimal solution and proves its optimality for all the 100 -asset and 200 -asset instances. The largest and more complex instances are those that clearly separate the outer approximation algorithms from the standalone Bonmin and the Cplex 12.1 solvers. The Bonmin solver cannot find any single integer feasible solution for any of the eight 1000-asset instances, while $O A A-B$ and $O A C-B$ find the optimal solution for $50 \%$ of the 1000 -asset instances and prove optimality in $50 \%$ of those cases. Similarly, OAA-C (resp., OAC-C) finds the optimal solution for $87.5 \%$ (resp., 100\%) of the 1000asset instances. Both algorithms prove optimality for $87.5 \%$ of the 1000 -asset instances, while Cplex finds the optimal solution for $62.5 \%$ and proves optimality for $37.5 \%$ of these instances. The results show without any ambiguity that the standalone mixed-integer nonlinear solvers Bonmin and Cplex are outperformed by the four outer approximation algorithms $O A A-B, O A C-B, O A A-C$ and $O A C-C$. Among these, $O A A-C$ and $O A C-C$ are the most performing ones in terms of solution quality.

The above conclusion is confirmed by Table 2 which reports the average optimality (AOG) and integrality (AIG) gaps obtained with the six methods. The notation $\infty$ indicates that no integer solution is found and that no gap value could be computed. The average optimality and integrality gaps obtained with $O A A-C$ and $O A C-C$ are systematically lower than those obtained with Cplex. The same conclusion applies when comparing OAA-B and OAC-B with Bonmin, with one exception. For the 700 -asset instances, Bonmin's average integrality gap is lower than the ones obtained with OAA-B and OAC-B. This is due to Bonmin finding a better lower bound (but weaker integer solution) than the ones obtained with OAA-B and OAC-B for the 700-4 instance.

| Number of Assets | OAA-B |  | OAC-B |  | Bonmin |  | OAA-C |  | OAC-C |  | Cplex |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG |
| 100 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.16 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 200 | 0.00 | 0.00 | 0.00 | 0.00 | 0.07 | 0.21 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| 500 | 0.02 | 0.11 | 0.02 | 0.22 | 0.06 | 0.27 | 0.00 | 0.07 | 0.00 | 0.07 | 0.05 | 0.18 |
| 700 | 0.02 | 0.39 | 0.00 | 0.34 | 0.04 | 0.33 | 0.01 | 0.17 | 0.01 | 0.18 | 0.02 | 0.26 |
| 1000 | 0.04 | 0.19 | 0.04 | 0.20 | $\infty$ | $\infty$ | 0.01 | 0.06 | 0.00 | 0.05 | 0.07 | 0.21 |

Computational time and scalability: We analyze the average (per instance family) time needed to show that the solution is optimal with OAA-B, OAC-B and Bonmin (Figure 8) and with OAA-C, OAC-C and Cplex (Figure 9). Figure 7 displays the average time needed by $O A A-B, O A C-B$ and Bonmin to find the optimal solution. These statistics are not reported for OAA-C, OAC-C and Cplex, since the Cplex 12.1 solver does not offer a convenient way to retrieve this information. The results displayed in Figures 7, 8 and 9 are based on the only instances for which optimality could be proven by each compared method.

Figures 7 and 8 show that there is no marked computational time differences between $O A A-B$ and $O A C-B$. To prove that the solution is optimal (resp., to find the optimal solution), OAA-B is on average 39.73 (resp., 54.51) seconds faster than Bonmin on the 100-asset instances and 1239.04 (resp., 1275.51) seconds faster than Bonmin on the 700 -asset instances. No speed comparison can be drawn for the 1000 -assets instances, since Bonmin is not able to find any feasible integer solution for any of those instances. With Bonmin, the computational times increase exponentially with the size of the problem, while they increase at a linear rhythm with OAA-B and OAC-B. As for the Bonmin solver, the outer approximation methods outperform the Cplex solver. To prove that the solution is optimal (Figure 9), OAA-C and OAC-C are on average 10.71 and 7.84 seconds faster than


Figure 7: Speed and Scalability with Bonmin: Average Time to Find the Optimal Solution


Figure 8: Speed and Scalability with Bonmin: Average Time to Show that the Solution is Optimal

Cplex on the 100-asset instances and 537.79 and 636.25 seconds faster than Cplex on the 1000 -asset instances.


Figure 9: Speed and Scalability with Cplex 12.1: Average Time to Show that the Solution is Optimal

It appears that the outer approximation algorithms need significantly less time than the standalone solvers Bonmin and Cplex to find optimal solutions as well as to show the optimality of the obtained solutions. The differences in computational times between the outer approximation algorithms and the solvers culminate for the largest 1000-asset instances. The computational times increase at a roughly linear rhythm with the four outer approximation algorithms which, in addition to being faster and finding solutions of better quality, scale much better than Bonmin and Cplex.

The speed of the outer approximation algorithms is due to their fast convergence and the very low number of iterations. For example, Tables 4 and 5 show that only one iteration is needed for 37 instances. Only three instances (100-5, 100-6 and 200-5) need two iterations (with OAA-B, OAC-B, OAA-C and OAC-C). This validates the criterion (see Section 3) used to select the assets and the associated binary variables on which we relax the integrality conditions. Clearly, our outer approximation algorithmic framework permits an easy recognition of the problem structure.

In order to differentiate the outer approximation algorithms with cuts ( $O A C-B, O A C-C$ ) from those without cuts ( $O A A-B, O A A-C$ ), we consider an extended set of instances and add, to those considered in Figures 7,8 and 9, some more complex instances that the standalone B\&B algorithms Bonmin and Cplex could not solve. The results displayed in Figure 10 and 11 concern all the instances that each compared algorithm can solve to optimality within one hour. It can be seen that $O A C-B$ is faster than $O A A-B$ for finding the optimal solution and for proving optimality. For the 100 -asset instances, the average time difference between OAA-B and OAC-B to find the optimal solution (resp., to prove that it is optimal) is 4.33 (resp., 2.42) seconds, whereas, for the 1000 -asset instances, the average time difference is 149.15 (resp., 162.03) seconds. To prove that the solution is optimal, the four algorithms need very similar solution times for instances including 500 or more assets. The two
outer approximation algorithms ( $O A C-B$ and $O A C-C$ ) including optimality cuts are slightly faster than ( $O A B-B$ and $O A B-C$ ) for the 700- and 1000 -asset instances.


Figure 10: Extended Set: Average Time to Find the Optimal Solution


Figure 11: Extended Set: Average Time to Show that the Solution is Optimal

### 4.2.3 Comparison of Solution Methods for Model with Buy-in Threshold Constraints

The analysis of the composition of the optimal index funds show that, in many instances, small positions are held by the investor. For 24 instances, the optimal index fund contains at least one position in which less than $1 \%$ of the investor's capital is invested (see Table 4). We shall now analyze the performance of the six methods for the solution of the more complex model ( $\mathbf{P B}$ ) incorporating buy-in threshold constraints (Section 2.5). The computational results were conducted by setting $w_{\min }$ equal to 0.01 in (23). Tables 6 and 7 (in Appendix) provide the same output and use the same notations as those used in Tables 4 and 5.

Solution quality: The dominance of the outer approximation algorithms OAA-B and OAC-B over Bonmin is even stronger than it is for the simpler model (P2) that does not include buy-in threshold constraints (see Figures 12 and 13). Besides not being able to find any integer solution for any of the 1000-asset instances (as for ( $\mathbf{P 2}$ )), Bonmin's performance weakens on smaller instances. Bonmin cannot find any integer solution for seven of the eight 700 -asset instances and for one of the 200 -asset instances. Moreover, Bonmin finds the optimal solution for only $62.5 \%$ (resp., $62.5 \%$ and $12.5 \%$ ) of the 200 -asset (resp., 500 - and 700 -asset) instances. The algorithms $O A A-B$ and $O A C-B$ cannot be differentiated in terms of the percentage of instances for which optimality is proven, but $O A C-B$ performs better than $O A C-B$ to find the optimal solution. In the same vein, the outer approximation algorithms OAA-C and OAC-C provide better results than Cplex (see Figures 14 and 15). Note that OAA-C and OAC-C find the optimal solutions for all but one and three instances, respectively, while Cplex does not find the optimal solution for 8 instances.


Figure 12: Solution Quality with Bonmin - Buy-In Threshold Model: Percentage of Instances when the Optimum is Found


Figure 13: Solution Quality with Bonmin - Buy-In Threshold Model: Percentage of Instances when Optimality is Proven


Figure 14: Solution Quality with Cplex 12.1 - Buy-In Threshold Model: Percentage of Instances when the Optimum is Found

Figure 15: Solution Quality with Cplex 12.1 - Buy-In Threshold Model: Percentage of Instances when Optimality is Proven

Computational time and scalability: Figure 16 displays the average time needed to find the optimal solution with $O A A-B, O A C-B$ and Bonmin. Figures 17 and 18 show the average time needed to prove that the solution is optimal. The graphs display only the $100-$, 200 - and 500 -asset instances, since the standalone solver Bonmin can find feasible integer solutions for only one 700-asset instance and for none of the 1000 -asset instances. The results displayed in Tables 16, 17 and 18 concern the instances for which optimality is proven by each algorithm. As for the model without buy-in threshold constraints, the four outer approximation algorithms are much faster than the standalone solvers Bonmin and Cplex. The time differentials between OAA-B and OAC-B on one hand and Bonmin on the other increase as the number of assets increases. The same conclusion prevails for the comparison between the two outer approximation algorithms OAA-C and OAC-C, and Cplex. This shows that the outer approximation algorithms scale better than the standalone solvers Bonmin and Cplex. We also observe that the four outer approximation algorithms converge fast (Tables 6 and 7). Only one iteration is needed for each algorithm on 36 of the 40 instances.


Figure 16: Scalability with Bonmin - Buy-In Threshold Model: Average Time to Find the Optimum


Figure 17: Scalability with Bonmin - Buy-In Threshold Model: Average Time to Prove that the Solution is Optimal


Figure 18: Scalability with Cplex - Buy-In Threshold Model: Average Time to Prove that the Solution is Optimal

Table 3 reports the average optimality (AOG) and integrality (AIG) gaps (in \%) with the six methods. The notation $\infty$ indicates that no representative gap measure could be obtained. The average optimality and integrality gaps obtained with the standalone solvers Bonmin and Cplex are systematically larger than the ones obtained with the outer approximation algorithms $O A A-B, O A C-B, O A A-C$ and $O A C-C$.

Table 3: Average Optimality and Integrality Gaps (in \%) - Buy-In Threshold Model

| Number of Assets | OAA-B |  | OAC-B |  | Bonmin |  | OAA-C |  | OAC-C |  |  | Cplex |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG | AOG | AIG |  |
|  | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |  |
| 200 | 0.00 | 0.00 | 0.00 | 0.00 | 0.08 | 0.30 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |  |
| 500 | 0.02 | 0.15 | 0.02 | 0.15 | 0.09 | 0.28 | 0.00 | 0.12 | 0.00 | 0.13 | 0.06 | 0.29 |  |
| 700 | 0.02 | 0.29 | 0.01 | 0.27 | $\infty$ | $\infty$ | 0.00 | 0.20 | 0.00 | 0.18 | 0.02 | 0.37 |  |
| 1000 | 0.05 | 0.19 | 0.03 | 0.17 | $\infty$ | $\infty$ | 0.00 | 0.19 | 0.01 | 0.22 | 0.02 | 0.26 |  |

The above analysis shows the versatility of the proposed outer approximation methods. The outer approximation algorithms OAA-B, OAC-B, OAA-C and OAC-C have been successfully applied to solve problem (PB) that includes buy-in threshold constraints. The problem (PB) with buy-in threshold constraints contains a larger number of integer constraints and poses additional combinatorial challenges than problems (P1) and (P2) for which the four algorithms were designed.

### 4.3 Analysis of Constructed Enhanced Index Funds

The proposed approach is intended to construct EIFs that perform well in-sample (i.e., on the data used for their construction), and, most importantly, out-of-sample, as one moves from the period used in their construction. For each problem instance, we use the index fund built on the basis of the in-sample period only (see Section 4.2) and verify how it performs on the one-year out-of-sample period.

Cross-Validation: The objective pursued in this study is to construct EIFs that closely track the S\&P 500 index while limiting the investors' overall risk exposure. To verify whether these objectives are attained during the out-of-sample period whose data were not used for the construction of the EIFs, we carry two types of crossvalidation experiments. The first one pertains to the tracking error and whether the EIFs and the S\&P 500 index share similar return patterns. The second experiment concerns the market risk exposure. More precisely, we check whether the EIFs' variance over the out-of-sample period is smaller than the maximum allowable variance level $v$ that is probabilistically enforced with the constraint (5).

First, in order to analyze how closely the constructed EIFs follow the S\&P 500 index, we use standard statistical hypothesis tests. The null hypothesis of the tests is that the monthly return tracking error between the S\&P 500 index and each constructed EIF is 0 . In other words, it means that on average there is no difference between the monthly returns of the S\&P 500 index and those of the considered EIF. Columns 2 to 11 in Tables 8 and 9 (for the model with buy-in constraints) report the $p$-value associated with each hypothesis test. The $p$-value is the probability of wrongly rejecting the null hypothesis if it is in fact true. A small $p$-value indicates that the null hypothesis is unlikely to be true: the smaller the $p$-value, the more reason one has to reject the null hypothesis. The $p$-value is compared to the actual significance level of the test which we set at the level ( $5 \%$ ) usually used. Clearly, the null hypothesis according to which the EIF closely tracks the S\&P 500 index is rejected if the $p$-value is strictly smaller than 0.05 .

For the in-sample period, for which we have 108 data points, we use the $t$-test. Columns named "IN" in Tables 8 and 9 show the $p$-value of the tests concerning the in-sample period. Not surprisingly, one can observe that the $p$-value is above 0.05 for each EIF and we do not reject the null hypothesis.

Next, we use the out-of-sample data points to cross-validate our approach and to check whether the constructed EIFs track closely the S\&P 500 index on the out-of-sample period. Since the out-of-sample period covers year 2006 and only contains 12 data points, it is not reasonable to employ the $t$-test. Hence, we use the Wilcoxon sign rank test for the out-of-sample period. It is a non-parametric test analogue to the $t$-test that does not require any assumptions about the distribution of the variable under consideration. Columns named "OUT" in Tables 8 and 9 report the $p$-value for each test concerning the out-of-sample period. It can be seen that the null hypothesis is rejected for only one instance (100-7). This confirms that our models permit the construction of EIFs that closely track the S\&P 500 index on (future) periods whose data were not used for the derivation of the model. The result is even more significant considering that we never rebalance the constructed EIFs. The EIFs are constructed on the basis on the 9 -year in-sample period. They track well the S\&P 500 index over that period and continue doing so over the next year, without any rebalancing. Also, note that the asset universe for the 40 problem instances was constructed completely randomly. For each problem instance, we selected at random a number $(n=100,200,500,700,1000)$ of assets out of the 2000 assets for which we had the necessary data for the period 1995-2006. In some problem instances, there was not a single asset included in the S\&P 500 index over the period 1995-2006.

Second, we evaluate whether the stochastic constraint (5) is actually effective for constructing risk-averse EIFs. Recall that we impose that the EIFs variance must be smaller than $v=0.005$ with probability $90 \%$. Columns 12-16 in Tables 8 and 9 (for the buy-in threshold model) report the variance of the EIFs over the out-of-sample period. Our risk averse-objective is clearly attained. The variance of each of the 80 EIFs is strictly smaller (i.e., the largest value is 0.0017 ) than the threshold of 0.005 .
Performance: After having shown that the constructed EIFs permit a close tracking of the S\&P 500 index while incurring a low overall variance, we shall now study the adjusted risk performance of the constructed EIFs. This is accomplished by comparing the Sharpe ratio [84] of the EIFs with the Sharpe ratio of the S\&P 500 index. The Sharpe ratio is calculated by subtracting the risk-free rate from the return of a portfolio and dividing the result by the standard deviation of the portfolio returns. It allows the verification of whether the portfolio's returns are due to good investment decisions or a result of excessive risk. An asset allocation is typically considered good if the (possibly high) returns it generates do not come with too much additional risk. The Sharpe ratio is a key metric to provide further insight into the performance of the constructed risk-averse EIFs. The Sharpe ratio of $97.5 \%$ of the constructed EIFs exceeds the Sharpe ratio of the S\&P 500 index fund.

## 5 Conclusion

Index fund allocation strategies have become increasingly popular due to the attractive return levels and the low transaction and rebalancing costs they incur. We propose a new model for the pursuit of a risk-averse enhanced indexation strategy and develop a new mathematical programming method based on the outer approximation concept. The method involves the early detection of the problem structure and develops a hierarchical organization of the computations with expanding sets of integer-restricted variables.

The proposed model constructs an index fund that mimics the return behavior of a benchmark market index. It implements a risk-averse partial replication strategy that limits the market risk incurred by the investor. The model takes into account the parameter estimation risk. We define asset returns and the return covariance terms as random variables, which implies that the variance of the index fund is itself a random variable. No probabilistic assumption is made on the asset returns that are driven by a set of factors whose return is stochastic. The risk-averse feature limits the return variability of the index fund. Our model includes a probabilistic constraint
that prevents the variance of the index fund return from exceeding a specified risk threshold with a large probability level. The motivation for this comes from the existing literature [7, 29, 49, 77] which reports that it is possible to design semi-active asset allocation strategies that reproduce very accurately the performance level of a benchmark, but at the cost of a sometimes high overall variance. Not only does the risk-averse feature limit the entailed market risk, but it also reduces the extent of a moral hazard problem related to the wage structure of fund managers [49]. Indeed, the remuneration of fund managers often includes a performance fee component, in which the salary is positively associated with the assumed risk [41, 49]. The risk constraint prevents the managers from constructing highly volatile EIFs and acts as a covenant.

The risk-averse EIF model takes the form of a stochastic integer programming problem that cannot be solved with any optimization solver in it original formulation. We propose two reformulation approaches to make the above problem computationally tractable. The first approach consists in the derivation of a new deterministic equivalent formulation for the stochastic constraint limiting the variance of the index fund. The second approach resides in the application of a block decomposition method that provides a much sparser representation of the VC matrix. This contributes to a significant decrease in computational times. The solution method involves the construction of increasingly tighter outer approximation problems obtained through the relaxation of the cardinality constraint and of the integrality restrictions on the binary variables associated with a carefully selected subset of assets. A variant of the proposed approach that involves the generation of an objective level optimality cut is also assessed.

We evaluate the performance of the outer approximation framework on 40 problem instances in which the EIF tracks the S\&P 500 index. The computational study shows that the outer approximation algorithm converges in a very small number ( 1 or 2 ) of iterations, and allows to solve very efficiently complex problem instances in which up to 1000 securities are considered for inclusion in the index fund. The strength of the obtained results is highlighted when linked to recent statements claiming that "traditional optimization techniques cannot deal appropriately with the discontinuities and the many local optima emerging from the introduction of explicit cardinality constraints" [66], or that exact methods "often cannot deal with the problem" [55] (i.e., index problem with asset universe comprising between 100 and 600 securities). Moreover, the algorithmic framework is general enough to allow the fast and optimal solution of problems including additional integer constraints, called buy-in threshold trading constraints, which prevent from holding small positions. The proposed approach outperforms the MINLP solvers Bonmin and Cplex 12.1 in terms of robustness (percentage of instances solved to optimality) and computational times, and scales much better. The conclusions apply to the two types of models (i.e., with or without buy-in threshold constraint) and are even more marked for the more combinatorially involved models including buy-in threshold constraints. The out-of-sample computational experiments show that the constructed EIFs track very closely the S\&P 500 index fund benchmark over periods whose data were not used for constructing the EIFs. More remarkable is that this result is obtained without any rebalancing operation over the entire (10-year) period. It is also shown that the EIFs expose the investors to a much smaller market risk than the one of the S\&P 500 index and that the variance of each of the EIF over the out-of-sample period is strictly smaller than the imposed threshold. More than $97 \%$ of the constructed EIFs generate a risk-adjusted return (i.e., Sharpe ratio) that is higher than the one of the S\&P 500 index.

Note that the same algorithmic technique could be applied if the fund manager is required to invest in a minimal number of assets. Future research will study the inclusion of transaction costs and of other trading constraints (turnover, etc.), as well as the inclusion of a constraint requiring the generation of a higher return level than the one of the benchmark (another form of enhanced indexation). Other algorithmic techniques involving, for exam-
ple, the polyhedral description of the continuous non-linear constraints or the derivation of lower objective level feasibility cuts will also be studied. Another objective is to consider non-normally distributed factor returns.

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Column 1 in Table 4 reports the name of the problem instance．Column 2 gives the optimal objective function value $Z^{*}$ ．Column 3 （resp．， 11 and 19 ）shows the best objective value $\bar{Z}$ found with OAA－B（resp．，OAC－B and Bonmin）．Column 4 （resp．， 13 and 20）reports the time to show that the solution is optimal with OAA－B（resp．，OAC－B and Bonmin）．For $\mathrm{OAC}-\mathrm{B}$ ，the time in column 13 is the sum of the times needed（i）to derive the optimality cut（column 12）and（ii）to show that the solution is optimal．When the optimality gap is not closed within one hour，the entry in the table reads＂$>3600$＂．Column 5 （resp．， 14 and 21）indicates the time to find the optimal solution with OAA－B （resp．，OAC－B and Bonmin）．For OAC－B，column 14 reports the sum of the times needed（i）to derive the optimality cut（column 12）and（ii）to find the optimal solution． When the optimal solution is not attained within one hour，the entry in the table reads＂＞3600＂．Column 6 （resp．，15 and 22）displays the optimality gap $O=\left(Z-Z^{*}\right) / \bar{Z}$ with OAA－B（resp．，OAC－B and Bonmin）．The symbol＂＊＂indicates that the optimal solution was found（ $O=0$ ）．Column 7 （resp．， 16 and 23 ）gives the best lower bound $\underline{Z}$ obtained with OAA－B（resp．，OAC－B and Bonmin）．We use＂＊＂to indicate that the optimality gap was closed and thus that $\underline{Z}=Z^{*}$ ；Column 8 （resp．， 17 and 24 ）reports the integrality gap $I=(\bar{Z}-\underline{Z}) / \bar{Z}$ with OAA－B（resp．，OAC－B and Bonmin）．The symbol ${ }^{*} *$＂indicates that the optimal solution was proven（ $I=0$ ）；Columns 9 and 18 provide the number of iterations with OAA－B and OAC－B．Column 10 reports the number of small（i．e．，$<0.01$ ）positions in the optimal index fund．The acronym NS indicates that no feasible integer solution was found．All computational times are in CPU seconds．Columns $6,8,15,17,22$ and 24 report the optimality and integrality gaps in percents．
Table 4：Computational Results for Problems without Buy－In Threshold Constraints with Bonmin

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Table 6: Computational Results for Problems with Buy-In Threshold Constraints with Bonmin

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| Number of AssetsData Set | 100 |  | 200 |  |  |  | 700 |  | 1000 |  | Outof:Sample Period Reurun Variance |  |  |  |  |
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|  | 3590 |  | .870 |  | 0.2110 |  | 0.1250 |  | 0.1410 |  | 0.00 |  |  |  |  |
| ${ }^{2}$ | ${ }_{0}^{0.4280}$ | ${ }^{0.1080}$ | 0.7510 | 0.1960 | ${ }^{0.3800}$ | ${ }^{0.5560}$ | 0.0790 | ${ }^{0.6100}$ | ${ }^{0.0590}$ | ${ }^{1.0000}$ | 0.0001 | 0.0002 | ${ }^{0.0003}$ | ${ }^{0.0005}$ | ${ }^{0.0008}$ |
|  |  | ${ }^{0}$ | ${ }^{0.0 .7600}$ |  | ${ }^{0.1390}$ | S470 | ${ }_{0}^{0.4700}$ | 0.400 | ${ }_{0}^{0.1000}$ |  | ${ }^{0.0002}$ | ${ }^{0} 0.0003$ |  | 2003 | ${ }_{0}^{0.0012}$ |
| 5 |  | ${ }_{0.2240}$ | 90 | ${ }_{0}^{0.1080}$ | ${ }_{0}^{0.3410}$ | 0.9060 | ${ }_{0}^{02650}$ | ${ }_{0}$. | ${ }_{0}^{0.0970}$ | ${ }_{0}^{0.4690}$ | ${ }_{0}^{0.0007}$ | ${ }_{0}^{0} 000022$ |  | 00012 | 007 |
| 6 | 0.4960 | 0.1080 | 0.9990 | 0.1470 | 0.2160 | 0.9690 | 0.0770 | 0.9060 | 0.0600 | 0.6100 | 0.0001 | 0.0002 |  | 006 | 011 |
| 7 | 0.7430 | 10 | ${ }^{0.8650}$ | ${ }^{0.5560}$ | ${ }^{0.2420}$ | ${ }^{0.8440}$ | 0.4410 | 0.4560 | 0.1650 | 0.7240 |  | 0.0002 | 0.0011 | 0.0004 | 0.0004 |
| 8 | 0.9120 | 0.1080 | 0.9 | 0.126 | 0.1870 | 0.7840 | 0.14 | 0.78 | 0.14 | 0.2 | 0.0 | 0.00 | 0.00 | 0.00 | 0.0004 |


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